

1. Questão 1

(a) Temos que:

$$\begin{aligned} f_{\mathbf{X}}(\mathbf{x}; \lambda) &= \exp \left\{ -n\lambda + \sum_{i=1}^n x_i \ln \lambda \right\} \left(\prod_{i=1}^n x_i \right) \mathbb{1}_A(\mathbf{x}) \\ &= \exp \{ c(\theta)t(\mathbf{x}) + d(\theta) \} h(\mathbf{x}) \quad (1) \\ &= \exp \{ \eta t(\mathbf{x}) + d_0(\eta) \} h(\mathbf{x}) \quad (2) \end{aligned}$$

Em que $c(\theta) = \eta = \ln \lambda$, $t(\mathbf{x}) = \sum_{i=1}^n x_i$, $d(\theta) = -n\lambda$, $\lambda = e^\eta$, $d_0(\eta) = -ne^\eta$.

Por (1) $\mathbf{X} \in FE_1(\lambda)$, T é suficiente, e, por (2), como $\Theta_\eta = (-\infty, 0)$ contem algum segmento de reta, T é completa e minimal, também.

Como $\mathcal{E}(T^*) = \mathcal{E}(T/n) = \lambda$, T^* é o ENVUM de λ .

(b) Pelo item anterior, temos que, $c(\cdot)$ é uma função crescente e, assim, \mathbf{X} tem RVMND em t . Portanto, um TUMP, para as hipóteses em questão, é dado por:

$$\phi(\mathbf{x}) = \begin{cases} 1, & \text{se } t(\mathbf{x}) > c \\ \gamma, & \text{se } t(\mathbf{x}) = c \\ 0, & \text{se } t(\mathbf{x}) < c \end{cases}$$

em que $t(\mathbf{x}) = \sum_{i=1}^n x_i$, $\alpha = P_{\lambda_0}(T > c) + \gamma P_{\lambda_0}(T = c)$, e $T \sim \text{Poisson}(n\lambda_0)$, se $\lambda = \lambda_0$.

(c) Temos, pelo item a), que:

$$L(\theta) \propto e^{-n\lambda} \lambda^{n\bar{x}} \mathbb{1}_{(0, \infty)}(\lambda)$$

Assim, a família conjugada corresponde à $\lambda \sim \text{gama}(a, b^{-1})$. Portanto,

$$\pi(\theta|\mathbf{x}) \propto e^{-(n+b)\lambda} \lambda^{n\bar{x}+a-1}$$

ou seja, $\theta|\mathbf{x} \sim \text{gama}(n\bar{x}+a, (n+b)^{-1})$.

(d) Pelo item anterior e o formulário, temos que: $2(n+b)\lambda|\mathbf{x} \sim \chi_{2a}^2$, em que $a^* = n\bar{x}+a$. Portanto:

$$P(q_1 \leq 2(n+b)\lambda \leq q_2 | \mathbf{x}) \leftrightarrow P\left(\frac{q_1}{2(n+b)} < \lambda < \frac{q_2}{2(n+b)} | \mathbf{x}\right)^\gamma$$

Logo, $IC_B(\lambda, \gamma) = \left[\frac{q_1}{2(n+b)}; \frac{q_2}{2(n+b)}\right]$, em que $P(X \leq q_1) = \frac{1-\gamma}{2}$ e $P(X \leq q_2) = \frac{1+\gamma}{2}$.

2. Questão 2

(a) Temos que:

$$\begin{aligned}\mathcal{E}(g(X)) &= \theta g(0) + 3\theta g(1) + (1 - 4\theta)g(2) = g(2) + (g(0) + 3g(1) - 4g(2))\theta \\ &\rightarrow \mathcal{E}(g(X)) = 0 \leftrightarrow g(2) + (g(0) + 3g(1) - 4g(2))\theta = 0 \leftrightarrow g(2) = 0; g(0) = -3g(1)\end{aligned}$$

portanto, não é completa

(b) Temos que;

$$L(x) = \frac{1}{-x + 3/2} < c \leftrightarrow -x > c^{-1} - 3/2 \leftrightarrow x < c^*$$

Assim, temos que um TUMP é dado por:

$$\phi(x) = \begin{cases} 1, & \text{se } x < c^*, \\ 0, & \text{cc} \end{cases}$$

em que $\alpha = P_{\theta_0}(X < c^*)$ e, se $\theta = 0$, temos que $X \sim U(0, 1)$ e, assim, $c^* = \alpha$

3. Questão 3

(a) Temos que

$$\mathcal{E}(T_a) = a\theta + (1 - a)c\mathcal{E}(S)$$

Mas, por outro lado, temos que: $Y = \frac{(n-1)S^2}{\theta^2} = \frac{2S^2}{2\theta^2/(n-1)} \sim \chi_{2(n-1)/2}^2$. Portanto,

$$(S^2 =) V = \frac{\theta^2}{n-1} Y \sim \text{gama}((n-1)/2, 2\theta^2/(n-1)) \equiv \text{gama}(b, d)$$

Portanto ($k > 0$, $b = (n-1)/2$, $d = 2\theta^2/(n-1)$)

$$\begin{aligned} \mathcal{E}(V^k) &= \int_0^\infty \frac{1}{\Gamma(b)d^b} e^{-v/d} v^{b+k-1} dv = \frac{\Gamma(a+k)b^{a+k}}{\Gamma(a)b^a} \\ &= \frac{\Gamma(b+k)d^k}{\Gamma(b)} \end{aligned}$$

Assim,

$$\mathcal{E}(V^{1/2}) = \frac{\Gamma(\frac{n-1}{2} + \frac{1}{2}) \frac{\theta\sqrt{2}}{\sqrt{n-1}}}{\Gamma(\frac{n-1}{2})} = \frac{\Gamma(\frac{n}{2}) \theta\sqrt{2}}{\sqrt{n-1} \Gamma(\frac{n-1}{2})}$$

Portanto, vem que:

$$\mathcal{E}(T_a) = a\theta + (1 - a)c\frac{1}{c}\theta = \theta$$

(b) Neste caso, temos que:

$$\mathcal{V}(T_a) = a^2 \frac{\theta^2}{n} + (1 - a)^2 c^2 \mathcal{V}(S)$$

mas

$$\mathcal{V}(S) = \mathcal{E}(S^2) - \mathcal{E}^2(S) = \theta^2 - \frac{\theta^2}{c^2} = \frac{\theta^2}{c^2}(c^2 - 1)$$

Logo:

$$\begin{aligned}\mathcal{V}(T_a) &= \frac{\theta^2}{n} (a^2 + n(c^2 - 1) - 2n(c^2 - 1)a + a^2n(c^2 - 1)) \\ &= \frac{\theta^2}{n} (n(c^2 - 1) - 2n(c^2 - 1)a + a^2[n(c^2 - 1) + 1]) \\ &= \frac{\theta^2}{n} (a^2b_1 + ab_2 + b_3)\end{aligned}$$

em que $b_1 = n(c^2 - 1) + 1$, $b_2 = -2n(c^2 - 1)$, $b_3 = n(c^2 - 1)$

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$$\Delta = b_2^2 - 4b_1b_3 = 4n^2(c^2 - 1)^2 - 4n^2(c^2 - 1)^2 - 4n(c^2 - 1)$$

Minimizar $g(a) = a^2b_1 + ab_2 + b_3$, $a' = -\frac{b_2}{2b_1} = 1$.

Assim, o estimar T_1 é o que apresenta a menor variância.

4. Questão 4

Temos que $F_X(x; \boldsymbol{\theta}) = \frac{x - \theta_1}{\theta_2 - \theta_1}$; $S_X(x; \boldsymbol{\theta}) = \frac{\theta_2 - x}{\theta_2 - \theta_1}$. Também,

$$f_{Y_1}(y; \theta_1) = n \frac{(\theta_2 - y)^{n-1}}{(\theta_2 - \theta_1)^n} \mathbb{1}_{(\theta_1, \theta_2)}(y); f_{Y_n}(y; \theta_2) = n \frac{(y - \theta_1)^{n-1}}{(\theta_2 - \theta_1)^n} \mathbb{1}_{(\theta_1, \theta_2)}(y)$$

(a) Temos que:

$$f_{\mathbf{X}}(\mathbf{x}; \theta_2) = \frac{1}{(\theta_2 - \theta_1)^n} \mathbb{1}_{(\theta_1, y_n)}(y_1) \mathbb{1}_{(\theta_1, \theta_2)}(y_n) = \frac{1}{(\theta_2 - \theta_1)^n} \mathbb{1}_{(\theta_1, y_n)}(y_1) \mathbb{1}_{(y_n, \infty)}(\theta_2)$$

Por outro lado, temos que:

$$\begin{aligned} \mathcal{E}(g(Y_n)) &= \int_{\theta_1}^{\theta_2} g(y) n \frac{(y - \theta_1)^{n-1}}{(\theta_2 - \theta_1)^n} dy = 0 \\ &\rightarrow \int_0^{\theta_2 - \theta_1} g(z + \theta_1) z^{n-1} dz = 0 \rightarrow g(\theta_2) = 0, \forall \theta_2 \in \Theta = (\theta_1, \infty) \end{aligned}$$

Logo, vale $\forall y_n \in (\theta_1, \theta_2)$. Logo Y_n também é completa. Por outro lado, temos que:

$$\begin{aligned} \mathcal{E}(Y_n) &= \int_{\theta_1}^{\theta_2} y n \frac{(y - \theta_1)^{n-1}}{(\theta_2 - \theta_1)^n} dy = \frac{n}{(\theta_2 - \theta_1)^n} \int_{\theta_1}^{\theta_2} y (y - \theta_1)^{n-1} dy \\ &= \frac{n}{(\theta_2 - \theta_1)^n} \int_0^{\theta_2 - \theta_1} (z + \theta_1) z^{n-1} dz = \frac{n}{(\theta_2 - \theta_1)^n} \left(\frac{(\theta_2 - \theta_1)^{n+1}}{n+1} + \theta_1 \frac{(\theta_2 - \theta_1)^n}{n} \right) \\ &= \frac{n(\theta_2 - \theta_1)}{n+1} + \theta_1 = \frac{n\theta_2 - n\theta_1 + n\theta_1 + \theta_1}{n+1} = \frac{n\theta_2 + \theta_1}{n+1} \\ &\rightarrow \mathcal{E} \left(\frac{(n+1)Y_n - \theta_1}{n} \right) = \theta_2 \end{aligned}$$

Logo $\frac{(n+1)Y_n - \theta_1}{n}$ é o ENVUM de θ_1 .

(b) Analogamente, temos que:

$$f_{\mathbf{X}}(\mathbf{x}; \theta_2) = \frac{1}{(\theta_2 - \theta_1)^n} \mathbb{1}_{(\theta_1, y_n)}(y_1) \mathbb{1}_{(\theta_1, \theta_2)}(y_n) = \frac{1}{(\theta_2 - \theta_1)^n} \mathbb{1}_{(-\infty, y_1)}(\theta_1) \mathbb{1}_{(y_1, \theta_2)}(y_n)$$

Por outro lado, temos que:

$$\begin{aligned}\mathcal{E}(g(Y_1)) &= \int_{\theta_1}^{\theta_2} g(y)n \frac{(\theta_2 - y)^{n-1}}{(\theta_2 - \theta_1)^n} dy = 0 \\ &\rightarrow \int_{\theta_2 - \theta_1}^0 g(\theta_2 - z)z^{n-1} dz = 0 \rightarrow g(\theta_1) = 0, \forall \theta_1 \in \Theta = (-\infty, \theta_2)\end{aligned}$$

Logo, vale $\forall y_1 \in (\theta_1, \theta_2)$. Logo Y_1 também é completa. Por outro lado, temos que:

$$\begin{aligned}\mathcal{E}(Y_1) &= \int_{\theta_1}^{\theta_2} yn \frac{(\theta_2 - y)^{n-1}}{(\theta_2 - \theta_1)^n} dy = \frac{n}{(\theta_2 - \theta_1)^n} \int_{\theta_1}^{\theta_2} y(\theta_2 - y)^{n-1} dy \\ &= -\frac{n}{(\theta_2 - \theta_1)^n} \int_{\theta_2 - \theta_1}^0 (\theta_2 - z)z^{n-1} dz = \frac{n}{(\theta_2 - \theta_1)^n} \int_0^{\theta_2 - \theta_1} (\theta_2 - z)z^{n-1} dz \\ &= \frac{n}{(\theta_2 - \theta_1)^n} \left(\frac{-(\theta_2 - \theta_1)^{n+1}}{n+1} + \theta_2 \frac{(\theta_2 - \theta_1)^n}{n} \right) \\ &= \frac{-n(\theta_2 - \theta_1)}{n+1} + \theta_2 \rightarrow \mathcal{E}(Y_1) = \frac{n\theta_1 + \theta_2}{n+1} \\ &\rightarrow \mathcal{E}\left(\frac{(n+1)Y_1 - \theta_2}{n}\right) = \theta_1\end{aligned}$$

Logo $\frac{(n+1)Y_1 - \theta_2}{n}$ é o ENVUM de θ_1 .

(c) Temos que

$$\begin{aligned}\mathcal{E}(Y_n + Y_1) &= \frac{n\theta_2 + \theta_1 + n\theta_1 + \theta_2}{n+1} = \theta_1 + \theta_2 \\ \mathcal{E}(Y_n - Y_1) &= \frac{(n-1)(\theta_2 - \theta_1)}{n+1} \rightarrow \mathcal{E}\left(\frac{n+1}{n-1}(Y_n - Y_1)\right) = \theta_2 - \theta_1\end{aligned}$$

Defina $V = Y_n + Y_1$ e $W = \frac{n+1}{n-1}Y_n - Y_1$. Assim, temos que:

$$\begin{aligned}\mathcal{E}(W + V) &= 2\theta_2 \rightarrow \mathcal{E}\left(\frac{W + V}{2}\right) = \theta_2 \\ \mathcal{E}(V - W) &= 2\theta_1 \rightarrow \mathcal{E}\left(\frac{V - W}{2}\right) = \theta_1\end{aligned}$$

Portanto, $\frac{W + V}{2}$ e $\frac{V - W}{2}$ são os ENVUM's de θ_2 e θ_1 , respectivamente.